

**CONTROLLING FOR OPTIMUM GROWTH
WITH TIME DEPENDENT RETURNS**

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IMA Preprint Series # 875

October 1991

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Abstract. Kelly (1956) gave a gambling interpretation of Shannon's notion of channel capacity, in the course of which a strategy now known as the Kelly Principle emerged: invest so as to maximize the expected logarithm of the return. Optimality of the Kelly Principle was considered from several points of view by Breiman (1961) who established rapid growth and least expected time to attain a large return. In this note, a new principle for achieving rapid growth is obtained for investments in which the returns are obtained over random amounts of time. This new principle is: invest so as to maximize the ratio of the expected logarithm of the return to the expected time to obtain it. The Kelly Principle corresponds to the case of a degenerate distribution for the time component. Breiman's interesting asymptotic formula comparing the time τ required to attain a large fortune C , in repeated play, versus the time τ^* required by the Kelly strategy is:

$$\lim_{C \rightarrow \infty} E(\tau - \tau^*) = \Sigma \left[1 - \frac{E \ln Y(\gamma_j, j)}{E \ln Y(\gamma^*, 1)} \right],$$

where $Y(\gamma_j, j)$ (resp. $Y(\gamma^*, j)$) is the random multiplicative factor by which the fortune at state j is transformed into the fortune at state $j + 1$, consequent to the predictably chosen strategic choice γ_j made for state j , asterisk denoting Kelly play. It is assumed that $\ln Y(\gamma^*, 1)$ is non-lattice in the above. We obtain Breiman's formula as the corresponding special case of a new formula:

$$\lim_{C \rightarrow \infty} E(\tau - \tau^*) = (E t(\gamma^*, 1)) \Sigma \left[\frac{E t(\gamma_j, j)}{E t(\gamma^*, 1)} - \frac{E \ln Y(\gamma_j, j)}{E \ln Y(\gamma^*, 1)} \right],$$

in which $t(\gamma_j, j)$ (resp. $t(\gamma^*, j)$) is the random time required for the j -th state of play to be completed, asterisk denoting ratio play. The new principle has obvious applications to gambling and investment and potential for generating fast converging algorithms for real-time estimation and prediction, in which one must decide between high-quality/infrequent estimate revisions versus low-quality/frequent estimate revisions.

Introduction. We begin with the model considered by Kelly (1956). Consider a finite probability experiment $\{p_1, \dots, p_k\}$ and non-negative numbers (multipliers) $\{m_1, \dots, m_k\}$. The idea is to bet on the outcome of this experiment with the understanding that a bet of f on outcome i will return $m_i f$ if outcome i occurs, and will otherwise return zero.

Whatever the amount of one's fortune f , and regardless of the state of play, bets may be placed on any or all of the outcomes, and some amount may be held back. Having a positive fortune f , one may bet on the outcomes $\{1, \dots, k\}$ by choosing amounts $\gamma_i \geq 0$, satisfying $\sum \gamma_i \leq 1$, and then proceeding to bet $f\gamma_i$ on outcome i . The return is

$$(\sum m_i \gamma_i (X = i) + (1 - \sum \gamma_i))f,$$

*Research partially supported by ONR Grants N00014-85-K-0150 and N00014-91-J-1087.

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where X denotes a random variable with $P(X = i) = p_i$, and the expression $(X = i)$ denotes the indicator random variable for the event $\{X = i\}$.

Following Kelly, assume an initial positive fortune F_0 and (after each play) reinvest the new fortune using the fixed betting fractions γ_i for the next (independent) play. Let F_n be the fortune after play n . Then

$$F_n = (\sum m_i \gamma_i (X_n = i) + (1 - \sum \gamma_i)) F_{n-1},$$

where $\{X_n\}$ are i.i.d. with $P(X_n = i) = p_i$. Write $F_n = Y(\gamma, n) F_{n-1}$ for the line above. As observed by Kelly, if $P(Y(\gamma, 1) = 0) = 0$, then with probability one

$$n^{-1} \ln (F_n/F_0) = \sum_{j=1}^n \ln Y(\gamma, j)/n \rightarrow E \ln Y(\gamma, 1).$$

and the limit is concave in γ . It follows that if $E \ln Y(\gamma^*, 1) > E \ln Y(\gamma, 1)$ then with probability one $F_n^* > F_n$ for all sufficiently large n . The Kelly Principle (broadly stated) is that, by choosing γ^* to maximize the expectation of the logarithm, the fortune F_n^* will eventually persist in higher returns than those consequent to any strategy which does not maximize the expected logarithm.

Time Dependent Returns. We now suppose that the factor $Y(\gamma, n)$ requires a random amount of time $t(\gamma, n)$ to be completed. The random vectors $(Y(\gamma, n), t(\gamma, n))$ are assumed to be i.i.d. for each γ , although Y and t need not be independent. The fortune $F_n = F_n(\gamma)$ is therefore available at time $T_n = \sum_1^n t(\gamma, j)$. For convenience, set $F_0 = 1$ and for each $C > 1$ define

$$(1) \quad \begin{aligned} N_C &= N_C(\gamma) = \min\{n : F_n \geq C\} \text{ (or } \infty), \\ \tau_C &= \tau_C(\gamma) = \sum_1^{N_C} t(\gamma, j). \end{aligned}$$

The following simple argument shows the correctness of the ratio principle.

PROPOSITION 1. *If $(E \ln Y(\gamma^*, 1))/E t(\gamma^*, 1) > (E \ln Y(\gamma, 1))/E t(\gamma, 1) > 0$ then, for every sufficiently large C , $E \tau_C^* < E \tau_C$.*

Proof. Wald's identity (Breiman, 1968) implies that the following are equivalent

$$(2) \quad \begin{aligned} E(\tau_C^* - \tau_C) &< 0 \\ (E N_C^*) E t(\gamma^*, 1) &< (E N_C) E t(\gamma, 1). \end{aligned}$$

For all sufficiently large C , the last line above is equivalent to $(E \ln Y(\gamma^*, 1))/E t(\gamma^*, 1) > (E \ln Y(\gamma, 1))/E t(\gamma, 1)$, since by standard renewal arguments, $\frac{E N_C^*}{E N_C} \rightarrow \frac{E \ln Y(\gamma, 1)}{E \ln Y(\gamma^*, 1)}$ as $C \rightarrow \infty$. \square

The next result generalizes Theorem 1 of Breiman (1961) to the case of time-dependent returns. It closely parallels the arguments before Theorem 1 of Breiman (1961).

PROPOSITION 2. If γ^* maximizes $(E \ln Y(\gamma, 1))/E t(\gamma, 1)$ for γ , $E \ln Y(\gamma^*, 1) > 0$, and $\ln Y(\gamma^*, 1)$ is non-lattice, then for every sequence of predictable strategies $\{\gamma_n\}$,

$$(3) \quad \lim_{C \rightarrow \infty} E(\tau_C^* - \tau_C)/(E t(\gamma^*, 1)) \\ = E \sum \left[\frac{E^{\mathcal{F}_{j-1}} t(\gamma_j, j)}{E t(\gamma^*, 1)} - \frac{E^{\mathcal{F}_{j-1}} \ln Y(\gamma_j, j)}{E \ln Y(\gamma^*, 1)} \right].$$

Proof. Define N_C^* and τ_C^* as in (1) for $\gamma = \gamma^*$; and N_C, τ_C for the sequence $\{\gamma_n\}$. Let $t_j = t(\gamma_j, j), t_j^* = t(\gamma^*, j), N = N_C, N^* = N_C^*, Y_j = Y(\gamma_j, j), Y_j^* = Y(\gamma^*, j), E^j = E^{\mathcal{F}_j}$ then

$$(4) \quad E \left[\sum_{t_j}^N - \sum_{t_j^*}^{N^*} \right] = E \sum (t_j - t_j^*) + E \sum t_j^* - E \sum t_j^*$$

By Wald's identity, (4) is equal to

$$(5) \quad = E \sum (t_j - t_j^*) + (E t^*)E(N - N^*) \\ = E \sum (E^{\mathcal{F}_{j-1}} t_j - E t^*) + \frac{E t^*}{E \ln Y^*} \left[E \sum (E \ln Y^* - E^{j-1} \ln Y_j) + D \right],$$

where $D = E \sum \ln Y_j - E \sum \ln Y_j^*$. As in Proposition 8 of Breiman (1961), on the event $\sum \ln Y_j - \sum \ln Y_j^* \rightarrow$ finite limit where $D \rightarrow 0$ a.s. . Elsewhere, the result (3) is trivial. Taking $C \rightarrow \infty$ in (4) gives

$$(6) \quad \lim_{C \rightarrow \infty} E \left[\sum_{t_j}^N - \sum_{t_j^*}^{N^*} \right] \\ = \lim_{C \rightarrow \infty} (E t^*) E \sum \left[\frac{E^{j-1} t_j}{E t^*} - \frac{E^{j-1} \ln Y_j}{E \ln Y^*} \right],$$

which is equal to the right side of (3) since the bracketed term on the right of (6) is non-negative. \square

Theorem 1 of Breiman (1961) is Proposition 2 above for the case $t_n \equiv t^* \equiv 1$.

Note. Breiman generalizes Kelly's model to allow for bets on events which are not disjoint. This generalization does not affect the above results or their proofs, which are valid for Breiman's model.

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